

PROMEMORIOR FRÂN P/STM

NR 8

HOW LARGE MUST THE SAMPLE SIZE BE? NOMINAL CONFIDENCE LEVELS VERSUS ACTUAL COVERAGE PROBABILITIES IN SIMPLE RANDOM SAMPLING AV JÖRGEN DALÉN

INLEDNING

TILL

Promemorior från P/STM / Statistiska centralbyrån. – Stockholm : Statistiska centralbyrån, 1978-1986. – Nr 1-24.

Efterföljare:

Promemorior från U/STM / Statistiska centralbyrån. – Stockholm : Statistiska centralbyrån, 1986. – Nr 25-28.

R & D report : research, methods, development, U/STM / Statistics Sweden. – Stockholm : Statistiska centralbyrån, 1987. – Nr 29-41.

R & D report : research, methods, development / Statistics Sweden. – Stockholm : Statistiska centralbyrån, 1988-2004. – Nr. 1988:1-2004:2.

Research and development : methodology reports from Statistics Sweden. – Stockholm : Statistiska centralbyrån. – 2006-. – Nr 2006:1-.

Promemorior från P/STM 1983:8. How large must the sample size be? Nominal confidence levels versus actual coverage probabilities in simple random sampling / Jörgen Dalén. Digitaliserad av Statistiska centralbyrån (SCB) 2016.



1983-02-14

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PREFACE

In survey sampling the prevailing estimation strategy for estimating a finite population parameter θ is based upon an (approximately) unbiased point estimator $\hat{\theta}$ and an (approximately) unbiased variance estimator $\hat{V}(\hat{\theta})$. Then the central limit theorem is referred to for an assumption that $\hat{\theta}$ is approximately normally distributed and it is stated that the interval

$$\hat{\theta} \pm 1.96 \times \sqrt{\hat{v}(\hat{\theta})}^7$$

covers the true value θ with a probability of approximately 95 %. Sometimes the value 1.96 is exchanged for the corresponding value taken from Student's t table with an appropriate number of degrees of freedom.

However, the central limit theorem does not say that this or that sample size is sufficient for the normal approximation to be reasonable in sampling from a particular finite population. For this reason much effort has been spent on the problem of convergence rates of sampling distributions to the normal distribution for simple random sampling from finite populations. But so far the attempts to reach exact theoretical results in this respect have more or less failed. This fact points at the necessity to seek other directions to approach this problem.

The non-applicability of the normal approximation is a problem that occurs frequently in sampling practice. One example is sampling from very skewed populations such as variables reported by enterprises (production, employment, investment, export, import). Another example is small area estimation where only a small number of observations form the basis of a certain figure in a table. For this reason a work has been initiated in the Statistical Methods Unit of Statistics Sweden concerning sampling from very skewed populations. One line of research has been to try to establish rules for when the normal approximation is applicable by making empirical studies of the distributional behaviour of estimators. In this paper such rules based on the population skewness are presented for the simple random sampling case.

The work is a part of a larger project within the Methods Unit concerning sampling and estimation led by Bengt Swensson. He, Carl-Erik Särndal and Jan Hagberg have also contributed to this report by giving helpful comments and suggestions.

1 Introduction

Erdös and Rényi (1959) and Hájek (1960) have developed conditions in finite populations for the sampling distribution of the sample mean to converge to normality. Höglund (1978) has made the following remainder term estimate (the formula is slightly manipulated algebraically to serve our purpose):

$$|F(x) - \Phi\left(\frac{x-n\mu}{\sigma\sqrt{n(1-f)}}\right)| \leq \frac{C}{\sqrt{n(1-f)}} \cdot G_{H}, \text{ where}$$

F is the distribution function of the sum of a sample of n elements among the N population elements $(X_1 \dots X_j \dots X_N)$, Φ the standard normal distribution function, μ is the population mean, σ is the population standard deviation, f = n/N,

C is an absolute constant and

$$G_{H} = \frac{\frac{1}{N} \sum_{j=1}^{N} |x_{j}-\mu|^{3}}{\sigma^{3}}$$

We notice that the deviation from normality is bounded by a term containing the factor G_{μ} , which may be considered to be a measure of the population skewness.

"For populations in which the principal deviation from normality consists of marked positive skewness", Cochran (1977) suggests the simple rule

n > 25
$$G_C^2$$
, where

$$G_C = \frac{\frac{1}{N} \cdot \frac{\sum_{j=1}^{N} (X_j - \mu)^3}{\sigma^3}$$

"This rule is designed so that a 95 % confidence probability statement will be wrong not more than 6 % of the time." We notice that the rule includes $G_c - a$ different measure of skewness (in fact the one most commonly used).

In this paper confidence intervals based on samples from the dichotomous population are systematically studied with respect to different population sizes, sample sizes and degrees of skewness. The reason for choosing the dichotomous population is that is is possible to obtain exact coverage probabilities for it. Also, the whole scale of different degrees of skewness is represented for a certain population size of the dichotomous population.

Exact probabilities are calculated for nominal 95 % confidence intervals based on the Student's t approximation to cover the true population mean. The t approximation is chosen instead of the normal one because it can be expected to give better approximations for small sample sizes and the 95 %-level because it is the most common one in practical survey sampling. Properties of these exact probabilities are demonstrated and simple rules of the Cochran type are proposed. The rules are formulated as $n > K_{\alpha} \cdot G^2$, where K_{α} is the constant required for obtaining at least α % probability for a supposed 95 % confidence interval to be true. The effects of using G_H as compared to G_C for G are studied. The rules are also found to work well for finite populations generated from some continuous parametric distributions.

2

The dichotomous population

For the dichotomous population studied the following notations are used:

Value	Number of elements	Number of elements
	in the population	in the sample
0	N - M	n – k
1	М	k
Total	N	n

The population has the following characteristics:

Population mean = μ = P Population variance = σ^2 = P - P² Skewness (Cochrans) = G_C = (1 - 2P)/(P - P²)^{0.5} Skewness (Höglunds) = G_H = (1 - 2P + 2P²)/(P - P²)^{0.5} where P = M/N.

Notice that $G_H = G_C + 2P^{1.5}/(1-P)^{0.5}$ so that $\lim_{P \to 0} (G_H - G_C) = 0$. Notice also that $G_c = 0$ and $G_H = 1$ when P = 0.5.

The sample has the following characteristics:

sample mean = $\overline{X} = k/n$ sample variance = $s^2 = (k-k^2/n)/(n-1)$

A supposed 95 % confidence interval for μ based on the sample outcome and the t-distribution would now be

$$\overline{X} - t_{0.975} \cdot s \cdot \sqrt{(1-f)/n} < \mu < \overline{X} + t_{0.975} \cdot s \cdot \sqrt{(1-f)/n}$$

(Since we are interested in how bad the t approximation could be at worst, the continuity correction is not used.)

Now, let I_k be the indicator for this confidence interval statement as a function of the sample outcome. That is:

 $I_{k} = \begin{cases} 1 \text{ for those } k \text{ when the confidence interval statement is true} \\ 0 \text{ for those } k \text{ when the confidence interval statement is false} \end{cases}$

The <u>actual coverage probability</u> (ACP) is now defined as the probability for a sample of a certain size n from our population to produce a true confidence interval statement. In mathematical notation this becomes

ACP(N, M, n) =
$$\sum_{k=0}^{n} p(k) \cdot I_{k}$$

where $p(k) = \frac{\binom{M}{k}\binom{N-M}{n-k}}{\binom{N}{n}}$

according to the hypergeometrical distribution.

Computer programs have been written, which compute these probabilities for various combinations of N, M and n. In table 1 one example of a result from such a program is shown.

There are many interesting features in this table. For example, there is not a steady increase in the ACP for increasing sample sizes. There are two reasons for this. One is due to the fact that the sampling distribution does not converge to the t-distribution when n approaches N. Actually, as pointed out by Plane and Gordon (1982), when n > N/2 the sampling distribution becomes more and more dissimilar to the normal distribution¹⁾.

The other reason for the jumps in the ACP as a function of sample size is understood when you look at the column "Outcomes where $I_k = 1$ ". There is, for example, a big downward jump from n=29 (ACP=95.6) to n=30 (ACP=82.6). This jump is due to the fact that the outcome k=1, which has a large probability to occur, for n=29 leads to a confidence interval, which just barely covers μ (the interval is from -0.0327 to 0.1016) but for n=30 the upper limit of the interval falls a little bit short of μ (it goes from -0.0313 to 0.0980).

In the example presented in table 1 you can also see that - due to the fluctuations of the ACP - you can never obtain a stable ACP-level of more than about 92 % by increasing the sample size (up to N/2).

1)

Plane and Gordon prove that the sampling distributions for n and N-n are mirror images of each other except for a scale change. However, as could be seen in Table 1, this does not mean that the ACP is the same for n and N-n. The scale factor is obviously important here. 3

A lot of tables like table 1 have been produced but for the sake of space they are not presented here. However, in table 2, we have calculated Cochran type constants leading to a certain guaranteed ACP. The procedure is as follows.

In deciding when a guaranteed ACP α is obtained we look for the smallest value of n (called n_{α}) such that ACP(n) $\geq \alpha$ for all $n \geq n_{\alpha}$. (For reasons mentioned above only $n \leq N/2$ are considered.) Then we calculate $K_{\alpha C} = n_{\alpha}/G_{C}^{2}$ and $K_{\alpha H} = n_{\alpha}/G_{H}^{2}$ whereupon we can obviously say that for all $n > K_{\alpha} \cdot G^{2}$ (indicated C or H) we have ACP(n) $\geq \alpha$.

For example look for a guaranteed ACP of 85 % in table 1. We see that $n_{0.85} =$ = 32, $K_{0.85C} = 32/2.67^2 = 4.5$ and $K_{0.85H} = 32/2.73^2 = 4.3$. These values could be found in the row in table 2 indicated N=300, M=30.

These calculations have been done for N=100 (100) 1 000, M=10, 20, 30, 40, 50, 75 and 100 (except for N=100 where M=5 (5) 50 is used), $2 \le n \le N/2$ and for $\alpha = 85$ %, 90 %, 93 %, 94 % and 94.5 %. The results are presented in table 2.

As we calculate $K_{\alpha} = n_{\alpha}/G^2$ and consider only $n \leq N/2$, we realize that for one particular population you can never obtain larger values of K_{α} than $N/2G^2$. The last two columns of table 2 and 3 present the value of this limit. Values of K_{α} close to the limit (80 % of it has been chosen as a simple rule) are put within brackets, since the constants obtained in these cases may be merely accidental from the point of view of all possible populations.

Some striking features in table 2 are

- Cochrans measure of skewness (G_C) is not at all suited to establishing criteria for populations with almost symmetrical distributions. Since $G_C = 0$ for symmetrical distributions $K_{\alpha C} \rightarrow \infty$ as $G_C \rightarrow 0$. When P > 0.25 it seems advisable not to use G_C any longer. G_H , on the other hand, is useful for all degrees of population skewness. When the skewness is large G_C and G_H are almost identical.
- Guaranteed ACPs of 94 % or more seem to be very difficult to obtain for these populations. $K_{\alpha C} = 25$, as suggested by Cochran, is clearly not enough. (We recognize, of course, that Cochran's rule is not intended to apply to dichotomous populations. For the equivalent case of proportions he suggests other criteria.) For $\alpha = 85$ % $K_{\alpha H} = 5$ seems to be sufficient in most cases and $K_{\alpha H} = 6$ in all cases. For $\alpha = 90$ % the corresponding figures would be $K_{\alpha H} = 10$ or 11. For $\alpha = 93$ % it is already more difficult to pass a judgement but it seems as if $K_{\alpha H} = 40$ would be sufficient.

4 Average ACPS

Instead of dealing with guaranteed ACPs we may define another, more "liberal", concept, called the average ACP, in the following way.

 β is an average ACP for a certain sample size n_0 in a certain population (M, N), if ACP(n_0) $\geq \beta$ and

$$\frac{\sum_{j=0}^{s} ACP(n_0+j)}{s+1} \ge \beta$$

for all values of s such that $0 \le s \le N/2 - n_0$.

In order to decide when an average ACP β is obtained we look for the smallest value of n_0^2 - called n_β^2 - such that the above conditions are valid. In the same way as before we then calculate $K_{\beta C}^2 = n_\beta^2/G_C^2$ and $K_{\beta H}^2 = n_\beta^2/G_H^2$. We then state that for $n > K_\beta^2 \cdot G^2$ (indicated C or H) there is an average ACP of at least β .

As an example we look for an 85 % average ACP in table 1. We see that ACP(18) > 85 %. There are only two ACPs smaller than 85 % for n > 18 (n = 30 and 31) and those are "averaged out" by the large ACPs surrounding them. We therefore certify that $n_{0.85} = 18$, $K_{0.85C} = 18/2.67^2 = 2.5$ and $K_{0.85H} = 18/2.73^2 = 2.4$. These values could be found in the row in table 3 indicated N=300, M=30.

What is the reason for bringing in the average ACP concept? It is obviously in order to even out the jumps in the ACP variation. In real life the exact population distribution is not known but you have an indication of the skewness from your sample. It is then intuitively reasonable to think of some kind of "expected ACP". Also, for larger and more diversified populations than the dichotomous one studied here, the jumps will be smaller and average ACP will be closer to the actual one.

The calculations based on the average ACP concept have been done for the same values of N, M, n and β (α) as was the case for the guaranteed ACPs. The results are presented in table 3.

What are the conclusions which could be drawn from this table?

- The relations between $\boldsymbol{G}_{\mbox{C}}$ and $\boldsymbol{G}_{\mbox{H}}$ is the same as in the preceding section.

- We generally get much lower constants than in the guaranteed ACP case. We also get useful results for the high β -levels. We see that a K_{BH} of 3-4 (2-3 for very skew populations) seems to be sufficient for 85 % average ACP. For 90 % K_{BH} = 5 seems to be enough, for 93 % 10-12, for 94 % around 20, while for 94.5 % it is more difficult to make a definite statement but K_{BH} = 35-40 seems to suffice.

5 Large populations

The populations studied above are comparatively small. N ranges from 100 up to 1 000. In certain instances there is a tendency of the constants K to rise with rising population size. It is therefore interesting to see what happens in much larger populations.

The rapid increase in computer time, however, preclude systematic investigations like those above for large populations. But it is possible to calculate ACPs for a limited selection of sample sizes. Tables 4-6 present some results of this kind.

In tables 4 and 5 some populations with different degrees of skewness are studied for some sample sizes. The population size is 10 000 in table 4 and 100 000 in table 5. For some sample sizes the corresponding value of the constant $(K_H = n/G_H^2)$ is also given. Because of the scattered sample sizes no definite conclusions can be drawn from these tables but there is nothing in them contradicting the above conclusions which are based on smaller population sizes.

Table 6 presents a population with a size of eight millions with

one percent ones and 99 percent zeroes. (This happens to be about the largest population of interest when sampling in Sweden.) ACPs are calculated for sample sizes from 50 to 6 000 with intervals of 50. $K_{\rm H} \approx 3$ for n=300, 5 for n=500, 11 for n=1 050, 20 for n=1 950 and 40 for n=3 900. Also in this case, the conclusions based on the small populations seem to be approximately valid.

6 Recommended values of K

The investigations presented so far could be summarized in the following way. In order to ensure at least a certain actual coverage probability (ACP), when sampling from a dichotomous finite population and using a t approximation for a 95 % confidence interval, the sample size n must be larger than K $\cdot G_{\rm H}^2$. The value of the constant K needed is a function of the desired ACP. It also depends upon the "average" or "guaranteed" interpretation of the ACP in about the following way:

	Value				
	85	90	93	94	94.5
Guaranteed	5-6	10-11	(40)	?	?
Average	2-4	5	10-12	20	40

We immediately realize that it is very difficult to formulate any rules with a "guaranteed" interpretation attached to them. But with the "average" interpretation this becomes possible and the K-levels get practical significance. The 94 %-rules is also close to the rule suggested by Cochran.

In the next section we investigate ACPs for other types of distributions based on K-levels of 3, 5, 11, 20 and 40.

7

Populations based on some continuous parametric distributions

In order to find out if the constants K could be applied to populations with structures entirely different from the dichotomous one, finite populations based on fixed percentiles of some continuous theoretical distributions were generated. These are:

I) The beta distribution with probability density function

$$f(x) = \frac{x^{a-1} (1-x)^{b-1}}{B(a,b)}$$
; $0 \le X \le 1$, $a > 0$, $b > 0$

where B(a,b) is the beta function.

$$\gamma$$
 = coefficient of skewness = $\frac{2(b-a)(a+b+1)^{1/2}}{(a+b+2)(ab)^{1/2}}$

II) The lognormal distribution with probability density function

$$f(x) = (\sigma x)^{-1} \cdot (2\pi)^{-1/2} \cdot \exp\left[-(\log x - \theta)^2/2\sigma^2\right]; x > 0, \sigma > 0$$

and

$$\gamma = \left[\exp(\sigma^2) + 2\right] \cdot \left[\exp(\sigma^2) - 1\right]^{1/2};$$

III) The power function with distribution function

$$F(x) = (x/\theta)^{c}; 0 < x < \theta, \theta > 0, c > 0$$

and

$$\gamma = \frac{2(1-c)(2+c)^{1/2}}{(3+c) \cdot c^{1/2}}$$

IV) The Weibull distribution with distribution function

$$F(\mathbf{x}) = 1 - \exp\left[-\theta \mathbf{x}^{c}\right]; \ \mathbf{x} \ge 0, \ \theta > 0, \ c > 0$$
$$\gamma = \frac{\Gamma(1+3/c) - 3\Gamma(1+2/c) \ \Gamma(1+1/c) + 2\Gamma^{3}(1+1/c)}{\left[\Gamma(1+2/c) - \Gamma^{2}(1+1/c)\right]^{3/2}}$$

where $\Gamma(\mathbf{x})$ is the gamma function.

In all cases γ stands for the coefficient of skewness corresponding to ${\tt G}_{\rm c}$ above, that is

$$\gamma = E(x-\mu)^3 / \left[E(x-\mu)^2\right]^{3/2}$$

The reference used for these distributions is Patel et al (1976).

A common feature of all these distributions is that for some value(-s) of the involved parameter(-s) γ can take on at least any value > 0.

For each of these four distributions six different finite populations have been generated with different degrees of skewness. The finite populations are all of size 500 and have been generated by taking the percentiles from 0.001 to 0.999 with intervals of 0.002.

For each population five different sample sizes have been chosen so that they correspond as closely as possible to the average K-levels above. This means that n has been chosen so that

$$n \ge K_{\beta H} \cdot G_{H}^{2} > n - 1$$

for values of K of 3, 5, 11, 20 and 40 respectively.

For every sample size 1 000 Monte-Carlo-simulations of a simple random sample without replacement have been made. For each sample the population mean has been estimated and a confidence interval based on the sample standard deviation and the t-distribution has been calculated. The number of cases when this interval covers the true population mean has been counted. This figure divided by 1 000 becomes our estimated actual coverage probability (EACP). EACP is of course stochastic in this case with a standard error of 0.7 % to 1.1 % when EACP ranges from 95 % to 85 %.

In table 7 - 10 the outcome of these simulations is presented in terms of the EACP for a certain combination of population and sample size. We see that in almost all cases the β -levels expected from the studies of the dichotomous population are surpassed, often by large margins. Only in five cases are the presupposed ACP-levels not obtained (those cases are indicated with an asterisk). The EACPs are in these cases 0.1 - 0.4 % below the expected level. One case is for K=11 (0.1 below), three cases are for K=20 (0.1-0.3 below) and one case is for K=40 (0.4 below). These small deviations in our results are not enough to cast any serious doubts on the usefulness of our proposed rules as a whole. The deviations may very well be entirely due to stochastic effects.

However, we must also admit that our rules for the higher ACP-levels $(\beta = 94 \text{ and } 94.5 \%)$ are not equally well founded compared with those for the lower levels $(\beta = 85 \text{ and } 90 \%)$. A more solid foundation for these levels would require a study of larger populations and sample sizes and also more replications for each case, since we are so close to 95 \% that the stochastic effect

of the simulations becomes an important disturbance.

8 Conclusions

Our empirical investigations have supported the following rule for the use of the Student's t approximation (and thereby also the normal approximation for reasonable sample sizes) in simple random sampling.

If you apply the t approximation for a 95 % confidence interval and the sample size

$$n > K_{\beta} \cdot G_{H}^{2}$$
,

you could be reasonably certain that the coverage probability is at least β . Compared with the rule suggested by Cochran a different measure of skewness is used. For populations with large positive skewness the two measures approximately coincide but for close-to-symmetrical populations it is necessary to use $G_{\rm H}$ instead of $G_{\rm C}$.

Five values of K have been derived, namely

$$K_{0.85} = 3,$$

$$K_{0.90} = 5,$$

$$K_{0.93} = 11,$$

$$K_{0.94} = 20 \text{ and}$$

$$K_{0.945} = 40.$$

 $K_{0.94}$ and $K_{0.945}$ remain not equally well founded compared with the others.

As a principal rule for allowing the t approximation in simple random sampling for a 95 % confidence interval, we propose

$$n > 20 \cdot G_{H}^{2}$$

If this rule is applied you should be able to count on being correct about 94 % of the time.

The values of K_{β} are not guaranteed in the sense that no instances could be found, where they are not sufficient. But in the average sense defined above they generally hold good and they are therefore sound to use in practical work.

However, one warning is necessary to issue. The rules above are based upon knowledge of the <u>population</u> skewness. It is not sufficiently known to what extent estimates of this skewness from the sample could replace it. For example, the sample skewness is a consistent but biased estimate of the population skewness. The bias could be considerable precisely in those cases where you would want to apply the sample skewness in the rules above. Work in this area is going on at the present time and results will be published in a forthcoming report.

What, then, are the theoretical reasons for these rules? In this area more research is undoubtedly needed in the future. Here, let us only point out the fact that the relevant convergence theorems and remainder term estimates are stated in terms of the standardized third moment of the population distribution. Therefore there is a reason to believe that two populations with the same degree of skewness should not differ radically with respect to the degree of correspondence between their sampling distributions and the t-distribution. 9

Cochran, William G., (1977), Sampling Techniques, third edition, Wiley

- Erdös, P & Rényi, A, (1959), On the central limit theorem for samples from a finite population. Publ. Math. Inst. Hung. Acad. Sci. 4, 49-57.
- Hájek, J, (1960), Limiting distributions in simple random sampling from a finite population. Publ. Math. Inst. Hung. Acad. Sci. 5, 361-374.
- Höglund, (1978), Sampling from a finite population. A remainder term estimate. Scand J Statist 5:69-71.
- Patel, J.K., Kapadia, C. H. & Owen, D.B., (1976), Handbook of statistical distributions, Marcel Dekker, Inc.
- Plane, D.R. & Gordon, K.R., (1982), A simple proof of the non-applicability of the central limit theorem to finite populations. The American statistician, Vol. 36, No. 3, Part 1.

C 'H' Outcomes (k)where $I_k=1$ Outcomes (k) where I_k=1 Sample size Sample size ACP ACP 68 69 23 18.060 1 2 93.625 1 -4 4 -

Table 1:	Actual	coverage	probability	(ACP) for	a dichotomous	population.
	N=300.	M=30. n=2	- 299. μ=0.1	• $G_c = 2.67$	G _u =2.73	

234567890123456789012345678901234567890123456789	$\begin{array}{l} 18.060\\ 0.923\\ 341.053713\\ 4472.537134\\ 45576659.4260237255\\ 6667777882857725531129\\ 663420023775552132255329\\ 6667777882857725531129\\ 66342002377553129\\ 66342002372553290233999999999999999999999999999999$	Noosoosoosoo a superior of the	 1233344444 4555555666666666777777777888 888888	6890 7712 77345677890 88888888890 9999999999999999999999	93.625 94.000 94.328 94.8512 94.8512 90.214 90.8879 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.3891 91.393 91.7351 91.4808 91.7351 93.4014 93.4014 94.4837 95.2629 91.7351 93.4014 95.4883 91.300	44444455555555555555556666666666666666	 222222333333333344444444455555555555555
34 4444444445555555555566666666666666666	91.880 93.101 93.667 94.170 94.611 94.694 95,323 87.932 91.149 91.748 90.479 91.748 91.749 91.748 91.749 91.748 91.729 92.7329 95.688 90.325 95.688 90.325 95.688 90.325 92.732 93.203	<u>ุงงงงงงงงพุทธุณ๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛๛</u>	 89 99999999999999999999999999999999999	$1023 \\ 1023 \\ 10067 $	92.683 93.110 93.502 93.1449 94.468970 94.468970 99.999 99.9999 99.99999 99.9999999999	777777777788888888888899999999999999	 16666666677777777777888888888889999

Table 1 (cont.)

Sample		Outcomes		Sample		Outcomes (
size	ACP	where I _k =	I	size	ACP	where I _k =1	
11111111111111111111111111111111111111	917695618166477226658866882029999999999999999999999999999999		19999999990000000000000000000000000000	890123456789012345678901221111112111112111111211111111111111	216998577749267&00726786647765310243922878961581487434091&04402194		ANNAARAAAAAAAAAAAAAAAAAAAAAAAAAAAAAAAA

		Outcon	nes (k)	
Sample		where	L = 1	,
size	ACP		<u>k</u>	
22222222222222222222222222222222222222	99999999999999999999999999999999999999	MANAANAANAANAA AAAAAAAAAAAAAAAAAAAAAAAA		уууууууууууууууууууууууууууууууууууу

	α	(%)								······			
	p	85		90		93		94		94.	5	$\frac{N}{2} \cdot \frac{1}{G^2}$	
		GC	G _H	GC	GH	G _C	G _H	GC	G _H	GC	G _H	^G с	G _H
N=100, M=5	0.050											2.9	2.9
, M=10	0.100	4.2	4.0	(6.3)	(6.0)	(6.8)	(6.4)	(6.9)	(6.6)	(7.0)	(6.7)	7.0	6.7
, M=15	0.150	3.1	2.8	10.2	9.0	(12.8)	(11.3)	(13.0)	(11.5)	-		13.0	11.5
, M=20	0.200	7.1	5.5	10.7	8.3	(21.3)	(16.6)	(21.8)	(17.0)	(22.2)	(17.3)	22.2	17.3
, M=25	0.250	5.3	3.4	15.0	9.6	(36.8)	(23.5)	(36.8)	(23.5)	(36.8)	(23.5)	37.5	24.0
, M=30	0.300	7.9	3.8	15.8	7.5	44.6	21.2	(59.1)	(28.1)	(64.3)	(30.6)	65.6	31.2
, M=35	0.350	12.6	3.8	15.2	4.6	48.0	14.6	-	-	-	-	126	38.3
, M=40	0.400	24.0	3.6	30.0	4.4	174	25.7	(280)	(42.6)	(280)	(42.6)	300	44.4
, M=45	0.450	99.0	3.9	124	4.9	371	14.6	(1 064)	(41.7)	(1 064)	(41.7)	1 238	48.5
, M=50	0.500	8	4	ω	8	∞	-	œ	-	œ	-	œ	50
N=200, M=10	0.050	3.5	3.5	(5.3)	(5.2)	(5.7)	(5.6)	(5.8)	(5.7)	(5.8)	(5.7)	5.9	5.8
, M=20	0.100	4.5	4.3	8.4	8.0	(13.9)	(13.3)	(14.1)	(13.4)	(14.1)	(13.4)	14.1	13.4
, M=30	0.150	5.5	4.8	8.6	7.6	21.3	18.8	(25.5)	(22.5)	(25.8)	(22.7)	26.0	23.0
, M=40	0.200	4.0	3.1	11.1	8.7	33.3	26.0	(43.1)	(33.6)	(43.6)	(33.9)	44.4	34.6
, M=50	0.250	5.3	3.4	15.0	9.6	42.0	26.9	56.3	36.0	(75.0)	(48.0)	75	48
, M=75	0.375	18.8	4.2	37.5	8.3	120	26.6	255	56.5	(371)	(82.2)	375	83
, M=100	0.500	∞	4	œ	11	œ	31	8	-	8	-	ω	100

<u>Table 2</u>: Cochran type constants (K_{α}) leading to a certain guaranteed ACP for all n > $K_{\alpha} \cdot G^{2}$ for nominal 95 % confidence intervals

Table 2 (cont.)

	$p \mid \frac{\alpha}{\alpha}$	(%)				<u>.</u>						N 1	
		85		90		93		94		94 .		$\frac{N}{2} \cdot \frac{1}{G^2}$	
······································		GC	G _H	GC	G _H	GC	с _н	^с с	G _H	GC	G _H	GC	G _H
N=300, M=10	0.033	3.3	3.3	(5.0)	(5.0)	(5.4)	(5.4)	(5.4)	(5.4)	(5.5)	(5.5)	5.5	5.5
, M=20	0.067	4.0	3.9	7.5	7.4	(12.3)	(12.0)	(12.3)	(12.1)	-	-	12.4	12.
, M=30	0.100	4.5	4.3	8.9	8.4	(20.5)	(19.5)	(21.0)	(19.9)	-	-	21.1	20.
, M=40	0.133	5.2	4.7	10.3	9.4	(31.2)	(28.3)	(31.4)	(28.5)	(31.4)	(28.5)	32.2	29.
, M=50	0.167	3.4	2.9	9.4	8.0	28.8	24.5	(46.9)	(39.9)	(46.9)	(39.9)	46.9	39.
, M=75	0.250	5.3	3.4	11.3	7.2	57.0	36.5	(112)	(71.5)	(112)	(71.5)	113	7
, M=100	0.333	10.0	3.6	22.0	7.9	80.0	28.8	(258)	(92.9)	(258)	(92.9)	300	10
N=400, M=10	0.025	3.2	3.3	(4.9)	(4.9)	(5.2)	(5.2)	(5.3)	(5.3)	(5.4)	(5.4)	5.4	5.
, M=20	0.050	3.7	3.7	8.5	8.4	(11.5)	(11.4)	(11.7)	(11.6)		-	11.7	11.
, M=30	0.075	4.1	4.0	8.1	7.9	(18.8)	(18.3)	(19.2)	(18.7)	(19.2)	(18.7)	19.2	18.
, M=40	0.100	4.5	4.3	8.9	8.4	(27.1)	(25.8)	(27.3)	(26.0)	(27.6)	(26.2)	28.1	26.
, M=50	0.125	5.1	4.7	9.9	9.1	(32.3)	(29.8)	(38.9)	(35.8)		-	38.9	35.
, M=75	0.188	3.9	3.2	10.5	8.5	34.7	28.0	(77.6)	(62.7)	-	-	78	63.
, M=100	0.250	5.3	3.4	11.3	7.2	39.8	25.4	(127)	(81.1)	(127)	(81.1)	150	9

Table 2 (cont.)

	\mathbf{p} α	(%)				· · · · ·							
	p	85		90		93		94		94.	5	$\frac{N}{2} \cdot \frac{1}{G^2}$	
		GC	G _H	GC	G _H	GC	G _H	GC	G _H	G _C	GH	GC	G _H
N=500, M=10	0.020	3.2	3.2	(4.8)	(4.8)	(5.1)	(5.1)	(5.2)	(5.2)	(5.3)	(5.3)	5.3	5.3
, M=20	0.040	3.6	3.6	8.3	8.2	(11.1)	(11.0)	(11.3)	(11.3)	-	-	11.3	11.3
, M=30	0.060	3.9	3.9	9.1	9.0	(17.8)	(17.6)	(18.1)	(17.8)	(18.1)	(17.8)	18.2	17.9
, M=40	0.080	4.3	4.2	10.0	9.7	(22.3)	(21.7)	(25.5)	(24.7)	(25.7)	(24.9)	26.1	25.3
, M=50	0.100	4.5	4.3	10.8	10.3	26.2	24.9	-	-	-	-	35.2	33.5
, M=75	0.150	5.7	5.1	11.2	9.9	33.1	29.2	(65.1)	(57.4)	(65.1)	(57.4)	65.1	57.4
, M=100	0.200	4.0	3.1	11.1	8.7	29.3	22.8	(105)	(82.0)	(111)	(86.2)	111	86.5
N=600, M=10	0.017	3.2	3.2	(4.8)	(4.8)	(5.1)	(5.1)	(5.2)	(5.2)	(5.3)	(5.3)	5.3	5.3
, M=20	0.033	3.5	3.5	(9.3)	(9.3)	(10.9)	(10.8)	(11.1)	(11.0)	-	-	11.1	11.0
, M=30	0.050	3.8	3.8	8.8	8.7	(17.3)	(17.1)	(17.4)	(17.2)	(17.5)	(17.3)	17.6	17.4
, M=40	0.067	4.1	4.0	9.5	9.3	(22.6)	(22.2)	(24.3)	(23.8)	(24.6)	(24.1)	24.9	24.4
, M=50	0.083	4.3	4.2	10.2	9.9	(23.2)	(22.5)	_	-	-	-	33.0	31.9
, M=75	0.125	5.1	4.7	10.1	9.3	36.8	33.9	(56.6)	(52.2)	(58.3)	(53.8)	58.3	53.8
, M=100	0.167	6.3	5.3	9.7	8.3	29.4	25.0	(90.9)	(77.5)	(90.9)	(77.5)	93.7	79.9

Table 2 (cont.)

	р	α	(%)	······································			·····				······································			
			85		90		93		94		94,5		$\frac{N}{2} \cdot \frac{1}{G^2}$	
			^G с	G _H	^G с	с _н	^с с	GH	GC	G _H	GC	G _H	^G с	G _H
N=700, M=10	0.0	14	3.1	3.1	(4.7)	(4.7)	(5.0)	(5.0)	(5.2)	(5.1)	(5.2)	(5.2)	5.2	5.2
, M=20	0.0	29	3.5	3.5	(9.2)	(9.2)	(10.7)	(10.7)	-	-		-	10.9	10.9
, M=30	0.0	43	3.7	3.7	10.0	9.9	(16.9)	(16.8)	(16.9)	(16.8)	(17.1)	(17.0)	17.2	17.0
, M=40	0.0	57	3.9	3.9	9.3	9.1	(21.9)	(21.6)	(23.6)	(23.2)	(23.9)	(23.6)	24.0	23.7
, M=50	0.0	71	4.2	4.1	8.1	7.9	22.1	21.6	-	-	-	-	31.6	30.9
, M=75	0.1	07	4.8	4.5	9.5	8.9	32.7	30.9	(46.6)	(44.0)	-	-	54.2	51.2
, M=100	0.1	43	5.5	4.9	8.6	7.7	28.6	25.6	67.7	60.6	(83.5)	(74.7)	84.0	75.2
N=800, M=10	0.0	13	3.1	3.1	(4.7)	(4.7)	(5.0)	(5.0)	(5.1)	(5.1)	(5.2)	(5.2)	5.2	5.2
, M=20	0.0	25	3.4	3.4	9.1	9.1	(10.6)	(10.6)	-	-	~	-	10.8	10.8
, M=30	0.0	38	3.6	3.6	9.8	9.8	(16.6)	(16.5)	(16.7)	(16.6)	(16.8)	(16.7)	16.9	16.8
, M=40	0.0	50	3.8	3.8	9.0	8.9	(21.4)	(21.2)	(23.1)	(22.8)	(23.3)	(23.1)	23.5	23.2
, M=50	0.0	63	4.0	3.9	9.4	9.3	24.1	23.7	-		-	-	30.6	30.1
, M=75	0.0	94	4.5	4.3	9.0	8.6	30.9	29.6	(51.5)	(49.3)	-	-	51.5	49.3
, M=100	0.1	25	5.1	4.7	10.1	9.3	30.1	27.8	59.5	54.8		-	77.8	71.7

Table 2 (cont.)

	$p \frac{\alpha}{\alpha}$	(%)										- > - 1	
	P	85		90		93		94		94.	5	$\frac{N}{2} \cdot \frac{1}{G^2}$	
		³ C	G _H	GC	G _H	G _C	G _H	GC	G _H	GC	G _H	GC	G _H
N=900, M=10	0.011	3.1	3.1	(4.7)	(4.7)	(5.0)	(5.0)	(5.1)	(5.1)	(5.2)	(5.2)	5.2	5.2
, M=20	0.022	3.4	3.4	(9.0)	(9.0)	(10.5)	(10.5)	-	-	-	-	10.7	10.7
, M=30	0.033	3.6	3.6	9.7	9.7	(16.4)	(16.3)	(16.5)	(16.4)	(16.6)	(16.5)	16.6	16.6
, M=40	0.044	3.8	3.8	8.8	8.7	(22.2)	(22.0)	(22.7)	(22.5)	(23.0)	(22.8)	23.0	22.8
, M=50	0.056	3.9	3.9	9.2	9.0	(24.8)	(24.5)	-	-	-	-	29.9	29.5
, M=75	0.083	4.3	4.2	8.7	8.4	28.3	27.4	-	_	-	-	49.5	47.9
, M=100	0.111	4.9	4.6	9.6	9.1	35.3	33.1	(73.0)	(68.6)	(73.5)	(69.0)	73.5	69.0
N=1 000, M=1	0 0.010	3.1	3.1	(4.7)	(4.7)	(5.0)	(5.0)	(5.1)	(5.1)	(5.2)	(5.2)	5.2	5.2
, M=2	20 0.020	3.4	3.4	(9.0)	(9.0)	(10.4)	(10.4)	-	-		-	10.6	10.6
, M=3	30 0.030	3.6	3.5	9.6	9.6	(16.2)	(16.1)	(16.3)	(16.2)	(16.5)	(16.4)	16.5	16.4
, M=4	0.040	3.7	3.7	8.6	8.6	(21.9)	(21.8)	(22.3)	(22.2)	(22.7)	(22.5)	22.7	22.5
, M=5	50 0.050	3.9	3.8	9.0	8.9	(26.9)	(26.6)	-	_	-	-	29.3	29.0
, M=7	75 0.075	4.2	4.1	10.0	9.7	35.9	35.0	-	-	-	-	48.0	46.8
, M=	100 0.100	4.6	4.4	9.3	8.8	32.1	30.5	(70.2)	(66.8)	(70.2)	(66.8)	70.3	66.9

	p β	(%)				• • • • • • • • •						Ň 1	
		85		90		93		94		94.		$\frac{N}{2} \cdot \frac{1}{G^2}$	
		^G С	GH	GC	G _H	GC	G _H	GC	G _H	GC	G _H	GC	G _H
N=100, M=5	0.050	1.8	1.8	2.2	2.2	-	-	-	-	-	-	2.9	2.9
, M=10	0.100	2.4	2.3	2.8	2.7	(6.8)	(6.4)	(6.9)	(6.6)	(7.0)	(6.7)	7.0	6.7
, M=15	0.150	3.1	2.8	3.6	3.2	6.5	5.8	(13.0)	(11.5)	-	-	13.0	11.5
, M=20	0.200	4.0	3.1	4.4	3.5	8.9	6.9	14.7	11.4	(22.2)	(17.3)	22.2	17.3
, M=25	0.250	5.3	3.4	6.0	3.8	7.5	4.8	19.5	12.5	(33.8)	(21.6)	37.5	24.0
, M=30	0.300	7.9	3.8	9.2	4.4	10.5	5.0	18.4	8.7	24.9	11.9	65.6	31.2
, M=35	0.350	12.6	3.8	15.2	4.6	17.7	5.4	17.7	5.4	50.6	15.3	126	38.3
, M=40	0.400	24.0	3.6	30.0	4.4	42.0	6.2	42.0	6.2	84.0	12.4	300	44.4
, M=45	0.450	99.0	3.9	124	4.9	124	4.9	149	5.8	446	17.5	1 238	48.5
, M=50	0.500	ω	4	ω	5	ω	5	œ	11	œ	11	œ	50
N=200, M=10	0.050	2.0	2.0	3.9	3.9	(5.7)	(5.6)	(5.8)	(5.7)	(5.8)	(5.7)	5.9	5.8
, M=20	0.100	2.5	2.4	5.1	4.8	(9.1)	(8.7)	(12.5)	(11.9)	(14.1)	(13.4)	14.1	13.4
, M=30	0.150	3.1	2.8	3.6	3.2	9.4	8.3	18.0	15.9	(25.8)	(22.7)	26.0	23.0
, M=40	0.200	4.0	3.1	4.9	3.8	8.9	6.9	20.9	16.3	31.1	24.2	44.4	34.6
, M=50	0.250	5.3	3.4	6.0	3.8	16.5	10.6	21.0	13.4	36.8	23.5	75	48
, M.=75	0.375	18.8	4.2	18.8	4.2	26.3	5.8	56.3	12.5	93.8	20.8	375	83.0
, M=100	0.500	8	4	œ	5	œ	5	œ	11	∞	14	œ	100

<u>Table 3</u>: Cochran type constants (K_{β}) leading to a certain average ACP for all $n > K_{\beta} G^2$ for nominal 95 % confidence intervals

Table 3 (cont.)

	p <u>β</u>	(%)										- <u></u>	
	P	85		90		93		94		94.	5	$\frac{N}{2} \cdot \frac{1}{G^2}$	
		G _C	G _H	GC	G _H	GC	G _H	GC	G _H	^G с	G _H	^G с	G _H
N=300, M=10	0.033	1.9	1.9	3.7	3.7	(5.4)	(5.4)	(5.4)	(5.4)	(5.5)	(5.5)	5.5	5.5
, M=20	0.067	2.2	2.2	4.5	4.4	8.2	8.0	(12.3)	(12.1)	-	~	12.4	12.2
, M=30	0.100	2.5	2.4	3.1	2.9	9.4	9.0	(18.0)	(17.1)	-	-	21.1	20.1
, M=40	0.133	3.0	2.7	3.4	3.1	9.0	8.2	19.1	17.4	24.9	22.7	32.2	29.3
, M=50	0.167	3.4	2.9	4.1	3.5	10.3	8.8	18.1	15.4	33.8	28.8	46.9	39.9
, M=75	0.250	5.3	3.4	6.8	4.3	12.0	7.7	25.5	16.3	25.5	16.3	113	72
, M=100	0.333	10.0	3.6	14.0	5.0	14.0	5.0	42.0	15.1	44.0	15.8	300	108
N=400, ^M =10	0.025	1.9	1.9	3.7	3.6	(5.2)	(5.2)	(5.3)	(5.3)	(5.4)	(5.4)	5.4	5.4
, M=20	0.050	2.1	2.1	4.2	4.2	9.0	8.9	(11.7)	(11.6)	-	-	11.7	11.6
, M=30	0.075	2.3	2.2	4.7	4.6	10.4	10.1	(16.5)	(16.1)	(19.2)	(18.7)	19.2	18.2
, M=40	0.100	2.5	2.4	5.2	5.0	9.7	9.2	15.1	14.3	(26.2)	(24.9)	28.1	26.8
, M=50	0.125	2.9	2.7	3.5	3.2	10.9	10.0	19.1	17.6	27.8	25.6	38.9	35.
, M=75	0.188	3.9	3.2	4.7	3.8	11.7	9.5	17.6	14.2	35.5	28.7	78	63.
, M=100	0.250	5.3	3.4	6.8	4.3	12.0	7.7	25.5	16.3	40.5	25.9	150	9

Table 3 (cont.)

	β	(%)											
	P	85		90		93		94		94.	5	$\frac{N}{2} \cdot \frac{1}{G^2}$	
		^G С	G _H	GC	G _H	GC	G _H	GC	GH	G _C	G _H	GC	G _H
N=500, M=10	0.020	1.8	1.8	3.6	3.6	(5.1)	(5.1)	(5.2)	(5.2)	(5.3)	(5.3)	5.3	5.3
, M=20	0.040	2.0	2.0	4.1	4.1	8.7	8.6	(11.3)	(11.3)	-	-	11.3	11.3
, M=30	0.060	2.2	2.2	4.5	4.4	9.7	9.5	15.5	15.3	(18.1)	(17.8)	18.2	17.9
, M=40	0.080	2.4	2.3	4.9	4.8	9.1	8.8	17.3	16.8	(25.7)	(24.9)	26.1	25.3
, M=50	0.100	2.5	2.4	5.3	5.1	11.5	11.0	20.4	19.4	(30.0)	(28.5)	35.2	33.5
, M=75	0.150	3.1	2.8	3.9	3.5	12.2	10.8	21.1	18.6	35.9	31.7	65.1	57.4
, M=100	0.200	4.0	3.1	4.9	3.8	12.4	9.7	21.8	17.0	31.1	24.2	111	86.5
N=600, M=10	0.017	1.8	1.8	3.6	3.6	(5.1)	(5.1)	(5.2)	(5.2)	(5.3)	(5.3)	5.3	5.3
, M=20	0.033	2.0	2.0	4.0	4.0	8.5	8.5	(11.1)	(11.0)	-	-	11.1	11.0
, M=30	0.050	2.2	2.2	4.4	4.4	9.3	9.2	(15.0)	(14.8)	(17.5)	(17.3)	17.6	17.4
, M=40	0.067	2.2	2.2	4.6	4.6	10.1	9.9	16.3	16.0	(24.6)	(24.1)	24.9	24.4
, M = 50	0.083	2.4	2.3	5.0	4.8	9.1	8.8	20.7	20.0	(31.0)	(30.0)	33.0	31.9
, M =75	0.125	2.9	2.7	3.5	3.2	10.9	10.0	17.3	16.0	33.8	31.2	58.3	53.8
, M=100	0.167	3.4	2.9	4.1	3.5	13.4	11.5	20.9	17.8	37.2	31.7	93.7	79.9

Table 3 (cont.)

	р <u> </u> В	(%)							<u></u> -			N 1	<u></u>
		85		90)	93		94		94.		$\frac{N}{2} \cdot \frac{1}{G^2}$	
		^G С	G _H	G _C	G _H	GC	G _Н	GC	G _H	GC	G _H	^с с	G _H
N=700, M=10	0.014	1.8	1.8	3.6	3.6	(5.0)	(5.0)	(5.2)	(5.1)	(5.2)	(5.1)	5.2	5.3
, M=20	0.029	2.0	2.0	4.0	4.0	8.4	8.3	-	-	-	-	10.9	10.9
, M=30	0.043	2.1	2.1	4.2	4.2	9.1	9.1	(14.5)	(14.4)	(17.1)	(17.0)	17.2	17.0
, ^M =40	0.057	2.2	2.2	4.5	4.4	9.8	9.7	17.0	16.8	(23.9)	(23.6)	24.0	23.
, ^M =50	0.071	2.4	2.3	4.7	4.6	10.5	10.2	18.5	18.1	(29.6)	(28.9)	31.6	30.
, ^M =75	0.107	2.6	2.5	3.3	3.1	12.2	11.6	19.5	18.4	33.2	31.3	54.2	51.
, ^M =100	0.143	3.1	2.8	3.6	3.2	9.4	8.4	18.7	16.8	33.1	29.6	84.0	75.
N=800, M=10	0.013	1.8	1.8	3.5	3.5	(5.0)	(5.0)	(5.1)	(5.1)	(5.2)	(5.2)	5.2	5.
, M=20	0.025	1.9	1.9	3.9	3.9	8.3	8.3	-	-	-	-	10.8	10.
, M=30	0.038	2.1	2.1	4.2	4.2	9.0	8.9	(15.5)	(15.4)	(16.8)	(16.7)	16.9	16.
, M=40	0.050	2.2	2.2	4.4	4.4	11.0	10.9	15.3	15.1	(23.3)	(23.1)	23.5	23.
, M=50	0.063	2.2	2.2	4.6	4.5	10.3	10.1	19.1	18.7	(28.6)	(28.1)	30.6	30.
, M=75	0.094	2.6	2.5	5.2	4.9	9.7	9.3	18.5	17.8	34.8	33.3	51.5	49.
, M=100	0.125	2.9	2.7	3.5	3.2	13.2	12.2	21.2	19.5	34.6	31.9	77.8	71.

Table 3 (cont.)

	pβ	(%)											
	F	85		90		93		94		94.	5	$\frac{N}{2} \cdot \frac{1}{G^2}$	
	ľ	^G C	G _H	GC	G _H	GC	G _H	^G с	G _H	GC	G _H	GC	G _H
N=900, M=10	0.011	1.8	1.8	3.5	3.5	(5.0)	(5.0)	(5.1)	(5.1)	(5.2)	(5.2)	5.2	5.2
, M=20	0.022	1.9	1.9	3.9	3.9	8.2	8.2	-	-	-	-	10.7	10.7
, M=30	0.033	2.0	2.0	4.1	4.1	8.9	8.8	(15.3)	(15.2)	(16.6)	(16.5)	16.6	16.6
, M=40	0.044	2.1	2.1	4.3	4.3	9.5	9.4	16.4	16.2	(23.0)	(22.8)	23.0	22.8
, M=50	0.056	2.2	2.2	4.5	4.5	10.0	9.9	17.4	17.2	(27.9)	(27.5)	29.9	29.5
, M=75	0.083	2.4	2.3	5.0	4.8	11.0	10.6	19.7	19.1	34.5	33.4	49.5	47.9
, M=100	0.111	2.8	2.6	3.3	3.1	12.4	11.7	22.2	20.9	40.7	38.2	73.5	69.0
N=1 000, M=1	0 0.010	1.8	1.8	3.5	3.5	(5.0)	(5.0)	(5.1)	(5.1)	(5.2)	(5.2)	5.2	5.2
, M=2	0.020	1.9	1.9	3.9	3.9	8.2	8.2	-	-	-	-	10.6	10.6
, M=3	0.030	2.0	2.0	4.1	4.0	8.8	8.8	(14.0)	(13.9)	(16.5)	(16.4)	16.5	16.4
, M=4	0 0.040	2.1	2.1	4.3	4.2	9.4	9.3	16.2	16.0	(22.7)	(22.5)	22.7	22.5
, M=5	0.050	2.2	2.2	4.4	4.4	9.8	9.7	17.1	16.9	(27.3)	(27.0)	29.3	29.0
, M=7	5 0.075	2.4	2.3	4.8	4.7	10.7	10.4	20.5	19.9	32.2	31.3	48.0	46.8
, M=1	000.100) 2.5	2.4	3.1	2.9	10.0	9.5	21.1	20.1	36.1	34.4	70.3	66.9

	M =																	
n =	5	(K _H)	10	(K _H)	50	(к _н)	100	(к _н)	200	(к _н)	500	(K _H)	1000	(к _н)	2000	(к _н)	5000	(K _H)
10	0.5		1.0		4.9		9.6		18.3		40.0		65.0		88.6	(3.5)	89.1	(10)
50	2.5		4.9		22.2		39.6		63.6		92.0	(2.9)	88.0	(6.7)	93.8	(17.3)	93.6	(50)
100	4.9		9.6		39.6		63.5		86.8	(2.1)	87.9	(5.8)	93.4	(13.4)	93.4	(34.6)	94.4	
300	14.1		26.3		78.2		80.4	(3.1)	93.5	(6.4)	92.9		94.1	(40.2)	95.0		94.7	
500	22.6		40.1		92.0	(2.6)	87.9	(5.2)	92.8	(10.6)	93.8	(29)	94.9		94.6		95.2	
1 000	41.0		65.1		88.6	(5.1)	93.4	(10.3)	93.6	(21.2)	94.9	(58)	95.2		94.9			
2 000	67.2		89.2	(2)	94.6	(10.2)	94.3	(20.6)	94.4		94.6		94.9		95.1			
3 000	83.2		84.9		94.8		94.1		94.6		94.8		94.9		95. 0			
4 000	92.2	(2)	95.2		93.9	(20.4)	94.5	(41.2)	94.9		94.4		94.7		95.0			
5 000	81.3		94.4	(5)	95.2		93.9		94.4		95.1		95.0		94.9			
Skewness (G _C)	44.7		31.6		14.0		9.8		6.9		4.1		2.7		1.5		0	
Skewness (G _H)			31.6		14.0		9.9		6.9		4.2		2.7		1.7		1	

Table 4: ACPs for some populations and sample sizes when N=10 000. Normal approximation is used.

	<u>M</u> =															
<u>n =</u>	10	(K _H)	50	(к _н)	100	(K _H)	500	(к _Н)	1000	(K _H)	2000	(K _H)	5000	(K _H) 10000	(K _H) 20000	(K _H)
10	0.1		0.5		1.0		4.9		9.6		18.3		40.0	65.0	88.6	(3.5)
50	0.5		2.5		4.9		22.2		39.5		63.5		92.0	(2.9) 87.9	(6.7) 93.8	(17.3)
100	1.0		4.9		9.5		39.4		63.4		86.7	(2.1)	87.8	(5.8) 93.	3 (13.4) 93.3	(34.6)
300	3.0		14.0		26.0		77.7		79.9	(3.1)	93.2	(6.4)	92.6	93.	3 (40.2) 94.7	
500	4.9		22.2		39.4		91.5	(2.6)	87.2	(5.2)	92.9	(10.6)	93.2	(29) 94.	94.9	
1 000	9.6		39.5		63.4		87.2	(5.1)	92.8	(10.3)	92.4	(21.2)	94.3	(58) 95.	94.8	
2 000	18.3		63.5		86.7	(2)	92.9	(10.2)	92.4	(20.6)	94.7	(42.4)	94.9	94.	9 94.8	
3 000	26.3		78.1		80.2		92.4		93.6		95.0		94.8	95.	95.1	
4 000	33.5		86.9	(2)	91.1		92.6	(20.4)	94.9	(41.2)	94.7		94.8	94.	7 94.9	
5 000	40.1		92.0		87.8	(5)	93.2		94.3		94.9		94.9	95.	95.0	
10 000	65.1		88.5	(5)	93.3	(10)	94.3	(51)	94.6							
20 000	89.2	(2)	94.6	(10)	94.2	(20)	95.4		95.2							
50 000	94.4	(5)	95.1	(25)	93.8	(50)	94.5		94.7							
Skewness (G	c) 100.0		44.7		31.6		14.0		9.8		6.9		4.1	2.	7 1.5	
Skewness (G	H) 100.0		44.7		31.6		14.0		9.9		6.9		4.2	2.	7 1.7	

Table 5: ACPs for some populations and sample sizes when N=100 000. Normal approximation is used.

	(r	ormal approxi	mation is	used.)	<u></u>			
		Outcomes where I _k =1			Outcomes where I _k =1			Outcomes where I _k =1
<u>n</u>	ACP	ĸ	<u>n</u>	ACP	K	<u>n</u>	ACP	K
50	39.5	1-4	1 750	92.0	12-27	3 450	94.6	25-47
100	63.3	1-5	1 800	93.6	12-28	3 500	93.8	26-48
150	77.8	1-6	1 850	94.3	12-28	3 550	94.8	26-49
200	86.5	1-7	1 900	92.9	13-29	3 600	93.4	27-49
250	91.5	1-7	1 950	94.3	13-30	3 650	94.4	27-50
300	79.9	2-8	2 000	94.9	13-30	3 700	94.8	27-50
350	86.2	2-9	2 050	93.6	14-31	3 750	94.1	28-51
400	90.7	2-10	2 100	94.8	14-32	3 800	95.0	28-52
450	93.3	2-10	2 150	93.0	15-32	3 850	93.8	29-52
500	87.1	3-11	2 200	94.3	15-33	3 900	94.7	29-53
550	90.8	3-12	2 250	94.8	15-33	3 950	95.0	29-53
600	93.0	3-12	2 300	93.7	16-34	4 000	94.4	30-54
650	88.3	4-13	2 350	94.9	16-35	4 050	94.7	30-54
700	91.4	4-14	2 400	93.2	17-35	4 100	94.1	31 - 55
750	93.7	4-15	2 450	94.4	17-36	4 150	95.0	31-56
800	89.4	5-15	2 500	94.9	17-36	4 200	93.8	32-56
850	92.0	5-16	2 550	93.9	18-37	4 250	94.7	32-57
900	94.1	5-17	2 600	94.4	18-37	4 300	95.0	32-57
950	90.4	6-17	2 650	93.4	19-38	4 350	94.4	33-58
1 000	92.7	6-18	2 700	94.5	19-39	4 400	94.7	33-58
1 050	93.9	6-18	2 750	95.0	19-39	4 450	94.1	34-59
1 100	91.3	7-19	2 800	94.1	20-40	4 500	95.0	34-60
1 150	93.3	7-20	2 850	94.6	20-40	4 550	93.8	35-60
1 200	94.4	7-20	2 900	93.7	21-41	4 600	94.7	35-61
1 250	92.2	8-21	2 950	94.7	21-42	4 650	95.0	35-61
1 300	94.0	8-22	3 000	93.2	22-42	4 700	94.4	36-62
1 350	94.8	8-22	3 050	94.3	22-43	4 750	94.7	36-62
1 400	93.0	9-23	3 100	94.7	22-43	4 800	94.2	37-63
1 450	94.5	9-24	3 150	93.9	23-44	4 850	95.0	37-64
1 500	92.0	10-24	3 200	94.9	23-45	4 900	93.9	38-64
1 550	93.7	10-25	3 250	93.5	24-45	4 950	94.7	38-65
1 600	94.5	10-25	3 300	94.5	24-46	5 000	95.0	38-65
1 650	92.9	11-26	3 350	94.9	24-46	5 050	94.5	39-66
1 700	94.3	11-27	3 400	94.2	25-47	5 100	94.8	39-66

Table 6: ACPs when N=8 000 000 and M=80 000 for some sample sizes $G_C \approx G_H = 9.85$. (Normal approximation is used.)

Table 6 (cont.)

n		ACP	Outcomes where I =1 k
5	150	94.3	40-67
	200	95.0	40-68
5	250	94.0	41-68
5	300	94.8	41-69
5	350	93.8	42-69
5	400	94.6	42-70
5	450	94.8	42-70
5	500	94.4	43-71
5	550	95.1	43-72
5	600	94.2	44-72
5	650	94.9	44-73
5	700	93.9	45-73
5	750	94.7	45-74
5	800	94.9	45-74
5	950	94.3	47-76
6	000	95.0	47-77

	- <u> </u>	b=1/a=		b=1/a	=1.756	b=1/a	=2.390	b=1/a	=4.083	b=1/a	=6.16		=8.509
К	∠5 ∠5	$G_{\rm H}=1.2$	299	G _H =1.	592	G _H =1.9	978	G _H =2.	946	G _H =3.	882	$G_{\rm H} = 4$.	714
<u>к</u>	<u>ک</u> (۵)	• <u>•</u> •••	n		'n		n		n		n		n
		G _C = (0	G _C =	1.000	$G_{\rm C} = 1$	1.598	_G =	2.767	G _C =	3.780	G _C =	4.647
3	85	94.2	6	90.2	8	90.9	12	89.8	27	88.5	46	89.3	67
5	90	94.6	9	92.6	13	90.9	20	90.9	44	91.7	76	91.6	112
11	93	94.8	19	93.7	28	93.3	44	94.1	96	94.2	166	94.6	245
20	94	95.1	34	95.5	51	93.9*	79	94.5	174				
40	94.5	94.7	68	95.0	102	94.6	157						

Table 7: Simulations from populations based on the beta distribution.

V	∂ ∠8(%)	$\frac{\sigma=0.1}{G_{H}=1.}$		σ=0.4 G _H =1.		σ=0.7 G _H =2.		σ=0.9 G _H =3.		$\sigma = 1.0$ $G_{\rm H} = 4.0$		$\sigma=1.3$ $G_{\rm H}=6.$	
K	(%)		n		n		n		n		n		n
		^G c ⁼⁰ .	220	^G _C =1.	25	^G c ⁼² .	577	^G C ⁼³ .	654	^G c ⁼⁴ .	533	^G C ⁼⁶ .	243
3	85	95.9	8	92.7	12	92.1	26	89.6	46	90.6	67	89.5	121
5	90	94.7	13	93.6	20	91.4	44	92.8	76	91.8	111	91.7	201
11	93	95.9	28	95.6	44	94.4	96	95.2	166	94.5	243		
20	94	94.8	51	95.5	79	93.9*	174						
40	94.5	95.4	102	96.2	158								

Table 8: Simulations from populations based on the lognormal distribution.

	(3 7)	$\frac{C=0.3}{G_{H}=1.}$		C=0.1 $G_{\rm H}=1.$		$\frac{C=0.09}{G_{\rm H}=2.9}$		C=0.0 $G_{\rm H}=3.3$	540 877	C=0.0 $G_{H}^{2}=4.$		C=0.0 G _H =6.	
K	∕B(%)		n	· · · · · · · · · · · · · · · · · · ·	n		n		n		n		n
		_G ⁼¹ .	117	^G C ⁼¹ .	702	G _C =2.8	329	^G c ⁼³ .	820	^G c ⁼⁴ .	663	^G c ⁼⁶ .	304
3	85	89.6	8	88.5	12	87.7	26	88.8	46	87.0	67	88.2	120
5	90	92.3	13	91.0	20	91.1	44	92.0	76	92.0	111	90.9	200
11	93	93.7	28	93.3	44	93.1	96	92.9*	166	94.9	243		
20	94	94.0	51	95.0	79	94.5	173						
40	94.5	94.7	102	95.1	157								

Table 9: Simulations from populations based on the power function distribution

	/3 B(%)	$\frac{C=2.15}{G_{H}}=1.585$		$\frac{C=1.24}{G_{H}=1.984}$		C=0.795 G _H =2.923		C=0.615 G _H =3.967		C=0.53 G _H =4.831		C=0.425 G _H =6.554	
К													
<u>к</u>			n		n		n		n		n		n
		G _C =0.528		G _C =1.406		G _C =2.645		G _C =3.809		G _C =4.724		G _C =6.499	
3	85	94.4	8	92.8	12	90.7	26	90.4	48	88.9	71	91.0	129
5	90	94.1	13	93.8	20	92.0	43	92.0	79	91.6	117	91.0	215
1	93	95.0	28	94.3	44	93.8	94	93.2	174				
0	94	94.5	51	93.7*	79	95.3	171						
0	94.5	95.2	101	94.1	158								

Table 10: Simulations from populations based on the Weibull distribution.

G-90693-30001-001-к-н

CATARINA ELFFORS

A/BF-S

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